Toward an Objective Characterization of an Anhedonic Phenotype: A Signal-Detection Approach

Diego A. Pizzagalli, Allison L. Jahn, and James P. O’Shea

Background: Difficulties in defining and characterizing phenotypes has hindered progress in psychiatric genetics and clinical neuroscience. Decreased approach-related behavior and anhedonia (lack of responsiveness to pleasure) are considered cardinal features of depression, but few studies have used laboratory-based measures to objectively characterize these constructs.

Methods: To assess hedonic capacity in relation to depressive, particularly anhedonic, symptoms, 62 participants completed a signal-detection task based on a differential reinforcement schedule. Anhedonia was operationalized as decreased reward responsiveness.

Results: Unequal frequency of reward between two correct responses produced a response bias (i.e., a systematic preference to identify the stimulus paired with the more frequent reward). Subjects with elevated depressive symptoms (Beck Depression Inventory scores ≥ 16) failed to show a response bias. Impaired reward responsiveness predicted higher anhedonic symptoms 1 month later, after controlling for general negative affectivity.

Conclusions: Impaired tendency to modulate behavior as a function of prior reinforcement might underline diminished hedonic capacity in depression. When applied to a clinical population, objective assessments of participants’ propensity to modulate behavior as a function of reward might provide a powerful tool for improving the phenotypic definition of depression and thus offer a reliable behavioral screening approach for neuroscience studies of depression.

Key Words: Depression, anhedonia, reward bias, reinforcers, appetitive behavior, phenotype

Recently, it has become increasingly evident that advancement in psychiatric genetics and clinical neuroscience is hindered by the heterogeneity of psychiatric disorders and by difficulties in defining and characterizing the phenotype under investigation. Therefore, failure to obtain reliable results regarding the etiology and pathophysiology of various psychiatric disorders might partially result from insufficient development of phenotypic characterization (Hasler et al 2004; Hyman 2002; Leboyer et al 1998). Leboyer et al, for example, suggested that identifying more homogeneous forms of diseases through a “candidate symptom approach” might yield better results than relying entirely on classic nosology. Similarly, Hasler et al (2004) proposed that a necessary step toward a better understanding of the genetic and neurobiological basis of depression will involve “dissecting” this complex clinical condition into key components yielding narrow definitions of psychopathologic endophenotypes (i.e., phenotypic traits or markers representing more direct expressions of underlying genes; Gottesman and Gould 2003).

Among promising endophenotypes, anhedonia (Hasler et al 2004)—the loss of pleasure or lack of reactivity to pleasurable stimuli—is one of the core symptoms of depression (American Psychiatric Association 1994) and has been considered a potential trait marker related to vulnerability to depression (Loas 1996; Meehl 1975). Costello (1972), for example, suggested that the impact of rewarding stimuli is markedly diminished in depression, and Meehl (1975) hypothesized that depressed subjects experience reward-related cues as less reinforcing due to a lowered hedonic capacity.

Various research traditions have provided empirical support for the hypothesis that depression is linked to fundamental deficits in an approach-related system promoting positive affect and motivated behavior. First, factor-analytic studies have generally found that low positive affect, particularly low engagement with the environment, uniquely differentiates depression from anxiety, whereas general negative affectivity represents a non-specific marker linked to both disorders (Watson and Clark 1984; Watson et al 1995). Second, electroencephalographic studies in depressed samples have generally reported dysfunctional brain activation in left and right prefrontal regions previously implicated in approach- and withdrawal-related affect, respectively (Davidson 1998; Gotlib et al 1998; Henriques and Davidson 1991; Pizzagalli et al 2002). Third, clinical observations and empirical findings have highlighted blunted engagement with the environment and diminished emotional response to pleasant cues in depression (Berenzbaum and Oltmanns 1992; Buchwald 1977; Gotlib 1982; Henriques et al 1994; Henriques and Davidson 2000; Lewinsohn and Graf 1973; Sloan et al 2001; Suslow et al 2001). This blunted engagement mirrors psychophysiological abnormalities in anhedonic subjects and melancholic patients (e.g. Bruder et al 1991; Fiorito and Simons 1994). In sum, substantial evidence suggests a critical role of an abnormal approach-related system in depression. Of clinical import, these dysfunctions might lead to diminished engagement in pleasurable activities and blunted responsiveness to reinforcing stimuli (reinforcers); this in turn could contribute to the generation, maintenance, or exacerbation of depressive symptoms, particularly symptoms of loss of pleasure, libido, appetite, and interest in the environment.

Considering the importance of anhedonia and abnormal approach-related behavior in depression, it is surprising to note that few studies have used laboratory-based measures to objectively characterize these constructs (e.g., Buchwald 1977; Henriques and Davidson 2000). Working toward the ultimate goal of a better phenotypic definition of depression, our aim in the present study was to assess the relationship between participants’ propensity to modulate behavior as a function of reward and depressive, particularly anhedonic, symptoms. To this end, a signal-detection task involving a differential reinforcement schedule with monetary reward was used to provide an objec-
tive, laboratory-based measure of hedonic capacity, which was operationalized as responsiveness to reinforcing stimuli (“reward responsiveness”).

In signal-detection tasks, participants are typically instructed to choose whether stimulus 1 or stimulus 2 was presented by making an appropriate response 1 or response 2 (McCarthy 1991). Participants’ performance can be decomposed into discriminability (sensitivity) and response bias. Discriminability measures the participants’ ability to differentiate between stimulus 1 and stimulus 2, whereas response bias reflects the participant’s general tendency to define an ambiguous stimulus as a target (i.e., the participant’s tendency to make one or the other response irrespective of which stimulus is presented). Critically, unequal frequency of reward between two types of correct responses typically produces a systematic preference for the response paired with the more frequent (or larger) reward (Macmillan and Creelman 1991; McCarthy 1991). In this study, the degree of response bias toward the more frequently reinforced alternative was used for operationalizing reward responsiveness, which can be understood as the extent to which behavior is modulated by reinforcement history. Note that this operationalization fits with the classical views that 1) a stimulus is rewarding if it positively reinforces behavior (i.e., it reliably increases the likelihood of a behavior) (Hull 1943; Rescorla and Wagner 1972); and 2) reinforcements play a critical role in the formation of associations between salient cues and internal rewarding events (Spanagel and Weiss 1999). Within this framework, we reasoned that blunted responses to reinforcements would provide a promising behavioral manifestation of hedonic hypofunctioning. If depression is associated with an impaired approach-related system and blunted responsiveness to reinforcing stimuli, we hypothesized that participants with elevated levels of depressive, particularly anhedonic, symptoms would show an impaired response bias toward the more frequently reinforced stimulus.

Methods and Materials

Participants

Informed written consent was obtained from 62 undergraduate students (23 men, 39 women; aged 20.13 ± 2.50 years [mean ± SD] recruited from introductory psychology courses at Harvard University. All participants were right-handed (Chapman and Chapman 1987), reported to be free of any past or present neurologic illness, and received either course credit or $5 for their participation. (Exploratory analyses revealed no differences in task performance, including response bias, between participants receiving course credit or $5 for their participation.) They were told they would have the opportunity to win money during the task. The study was approved by the Committee on the Use of Human Subjects at Harvard University.

Task and Procedure

Participants were given verbal instructions about the task and told that the aim of the task was to win as much money as possible (actual money was shown to the participants to increase the credibility of the experiment). Participants then completed the Chapman Handedness Inventory (Chapman and Chapman 1987). Subsequently, written instructions about the task, a short demonstration, and two practice trials were presented. The task, which was adapted from Tripp and Alsop (1999), was presented on a Macintosh PowerMac G4 (866MHz) computer (Apple Computer, Cupertino, California) running Psyscope software (Cohen et al 1993) and lasted approximately 30 min. The task consisted of 300 trials, divided into 3 blocks of 100 trials, with blocks separated by a 30-sec break. Each trial started with the presentation of an asterisk for 500 msec in the middle of the screen that served as a fixation point (Figure 1). After 500 msec, a mouthless cartoon face was presented in the center of the screen. After a delay of 500 msec, either a short mouth (11.5 mm) or a long mouth (13 mm) was presented for 100 msec. The face without the mouth remained on the screen until a key response was made. Participants were asked to identify which type of mouth was presented by pressing either the “z” or the “/” key (counterbalanced across subjects). For each block, the long and short mouths were presented equally often in a pseudorandomized sequence, with the constraint that no more than three instances of the same stimulus were presented consecutively.

An asymmetrical reinforcer ratio (the relative number of reinforcers received after a given correct response vs. another correct response) was used because both animal and human research have shown that this is the critical variable for producing a response bias (Johnstone and Alsop 2000; McCarthy and Davison 1979). In this task, the only type of feedback provided was reward feedback for correct responses. Subjects were specifically instructed that not all correct responses would be receiving reward feedback. For each block, 40 correct trials were followed by reward feedback (“Correct! You won 5 cents”), presented for 1750 msec immediately after the correct response. For half of the participants, correct identification of the short mouth was associated with three times more positive feedback (30 of 40) than correct identification of the long mouth (10 of 40). For the other half of the participants, the contingencies were reversed. A controlled reinforcer procedure was used so that reward feedback was given according to a pseudorandom schedule that determined which specific trials were to be rewarded for correct identifications (Johnstone and Alsop 2000; McCarthy and Davison 1979). If a subject failed to make a correct identification in a trial for which reward feedback was due according to the schedule, the feedback was delayed until the next correct identification of the same stimulus type. For the entire task, participants “earned” approximately $6.

After the computer task, participants completed several self-report measures of affect and mood. The trait form of the Positive and Negative Affect Scales (PANAS; Watson et al 1988) was used to assess two fundamental dimensions of dispositional affect: positive and negative affect. The Beck Depression Inventory-II (BDI-II), a reliable and well-validated self-report instrument of depressive symptomatology (Beck et al 1996), was used to assess...
levels of depressive symptoms. Levels of trait anxiety were assessed with the Spielberger Trait Anxiety Inventory (STAI; Spielberger et al 1970). Finally, anhedonia was assessed by 1) the Revised Physical and Social Anhedonia Scales (Chapman and Chapman, unpublished data; Chapman et al 1976, 1980); and 2) a total score on BDI items associated with anhedonic symptoms (BDI anhedonic subscore; Cochrane α = .60): loss of pleasure (item 4), loss of interest (item 12), loss of energy (item 15), and loss of interest in sex (item 21). Finally, in light of our interest in melancholic depression (Pizzagalli et al 2004)—a subtype of major depressive disorder characterized by pervasive anhedonia (Rush and Weissenerburger 1994)—a total “melancholic” symptom score was computed by summing scores of BDI items that map onto the DSM-IV criteria for melancholia (BDI melancholic subscore; Cochrane α = .45): loss of pleasure (item 4), guilty feelings (item 5), agitation (item 11), loss of interest (item 12), early morning awakening (item 16b), and loss of interest in sex (item 21) (Pizzagalli et al 2004).

Table 1 lists the zero-order (Pearson) correlations between the various self-report measures of affect and mood.

To assess the test–retest reliability of performance in the signal-detection task and to test whether low response bias at the initial behavioral session (Time 1) predicted subsequent self-report measures of affect, a randomly selected subsample of the participants was recontacted at a later point and invited to participate in an identical session (Time 2). At Time 2, the BDI and the trait version of the PANAS were administered again, along with the Mood and Anxiety Symptom Questionnaire (MASQ; Watson et al 1995), which assesses anhedonic depression, anxious arousal, and symptoms of general distress and thus provided an independent assessment of anhedonic symptoms. Twenty-five participants performed the signal-detection task a second time (on average, 38.28 days after the first session; range, 21–120 days).

Data Collection and Reduction

Task performance was analyzed with respect to discriminability, reaction time (RT), and response bias. Whereas discriminability and RT provide information about overall task performance (e.g. attention, response speed), response bias provides an index of the participants’ systematic preference for the response paired with the more frequent reward, and was thus the variable of main interest. Discriminability and response bias were derived from the behavioral model of signal detection (e.g., McCarthy and Davison 1979; Tripp and Alsop 1999). Specifically, discriminability was computed as:

$$\log d' = \frac{1}{2} \log \left( \frac{\text{Long}_{\text{correct}} \ast \text{Short}_{\text{correct}}}{\text{Long}_{\text{incorrect}} \ast \text{Short}_{\text{incorrect}}} \right)$$

where Long_{correct}, for example, is the number of correct responses after presentation of the long mouth, Short_{incorrect} is the number of incorrect responses after presentation of the short mouth, and so on. Discriminability is mainly affected by physical differences between the two target stimuli and thus provides information about task difficulty. Response bias was computed as:

$$\log b = \frac{1}{2} \log \left( \frac{\text{Rich}_{\text{correct}} \ast \text{Learn}_{\text{correct}}}{\text{Rich}_{\text{incorrect}} \ast \text{Learn}_{\text{correct}}} \right)$$

assuming that correct identification of the long mouth in this example was rewarded more frequently. High response bias scores are produced by high numbers of correct responses to the stimulus associated with more frequent reward (“rich condition”) and high numbers of incorrect response to the stimulus associated with less frequent reward (“lean condition”), which would increase the numerator and decrease the denominator. $^1$ Log d is similar to traditional signal-detection measures, such as $d'$ (Johnstone and Alsop 2000). Likewise, log b is functionally equivalent to the criterion location, c, or the choice theory measure, log (b) (Luce 1963; Green and Swets 1966).

For all analyses, trials with RT shorter than 150 msec or longer than 2500 msec were excluded (overall, 59% of the trials). Furthermore, for each subject, trials with RT (after natural log transformation) falling outside the range of mean ± 3 SD were considered outliers. Overall, an additional 1.11% of trials were excluded.

$^1$Formula 2 can be generalized as:

$$\log b = \frac{1}{2} \log \left( \frac{\text{Rich}_{\text{correct}} \ast \text{Learn}_{\text{correct}}}{\text{Rich}_{\text{incorrect}} \ast \text{Learn}_{\text{correct}}} \right)$$
Table 2. Demographic and Self-Report Data for High (N=15) and Low (n=21) BDI Subjects

<table>
<thead>
<tr>
<th></th>
<th>Low BDI</th>
<th>High BDI</th>
<th>t</th>
<th>p</th>
</tr>
</thead>
<tbody>
<tr>
<td>n</td>
<td>21</td>
<td>15</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Female/Male</td>
<td>14/7</td>
<td>10/5</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age (y)</td>
<td>20.76 ± 3.94</td>
<td>19.87 ± 1.41</td>
<td>.84</td>
<td>.40</td>
</tr>
<tr>
<td>BDI</td>
<td>3.62 ± 2.27</td>
<td>21.33 ± 4.82</td>
<td>−14.34</td>
<td>.001</td>
</tr>
<tr>
<td>STAIa</td>
<td>33.92 ± 7.47</td>
<td>48.44 ± 6.69</td>
<td>−5.77</td>
<td>.001</td>
</tr>
<tr>
<td>PANAS-PA (Trait)</td>
<td>35.15 ± 7.29</td>
<td>32.67 ± 5.75</td>
<td>1.09</td>
<td>.28</td>
</tr>
<tr>
<td>PANAS-NA (Trait)</td>
<td>13.65 ± 4.39</td>
<td>20.47 ± 6.94</td>
<td>−3.55</td>
<td>.002</td>
</tr>
<tr>
<td>Social Anhedonia</td>
<td>5.76 ± 4.68</td>
<td>7.00 ± 5.64</td>
<td>−.72</td>
<td>.48</td>
</tr>
<tr>
<td>Physical Anhedonia</td>
<td>10.29 ± 7.08</td>
<td>12.00 ± 8.04</td>
<td>−.68</td>
<td>.50</td>
</tr>
</tbody>
</table>

Data are presented as n or Mean ± SD.

Notes: BDI, Beck Depression Inventory-II (Beck et al 1996); STAI, Spielberger Trait Anxiety Inventory (Trait form; Spielberger et al 1970); PANAS, Positive and Negative Affect Scales (PA, positive affect; NA, negative affect; Watson et al 1988). Revised Physical and Social Anhedonia Scales (Chapman and Chapman, unpublished data; Chapman et al 1976; Chapman et al 1980).

*Low BDI (n = 12) and high BDI (n = 10).

Statistical Analyses

Exploratory analyses revealed no differences in behavioral performance between female and male participants; consequently, analyses were performed on the entire sample. For accuracy (percent correct response) and RTs, analyses of variance (ANOVA) with Block (1, 2, 3) and Condition (rich, lean) as repeated measures were performed. For response bias and discriminability, the ANOVAs included Block only. Significant ANOVA effects were further explored with post hoc Newman-Keuls tests. When applicable, the Greenhouse-Geisser correction was used, and epsilon (ε) values are reported. Partial eta-squared (η²) values are reported as measures of effect sizes.

To assess whether elevated levels of depressive symptoms modulated the effects, participants with low BDI (range, 0–6, n = 21) and high BDI (range, 16–35, n = 15) scores were compared with Group as an additional between-subjects factor. According to the BDI-II manual, a score of >16 reflects mild depression, and a study combining the BDI-II and a structured clinical interview in a student sample has shown that a score of 16 is an effective cutoff for determining depressed mood (Sprintke et al 2002). High BDI (score 21.33 ± 4.82) and low BDI (score 3.62 ± 2.27) subjects did not differ with respect to age or gender ratio (Table 2). The two groups had comparable percentages of trials excluded from the analyses because their RT was too short (>150 msec), too long (>2500 msec), or was considered an outlier response (exceeding mean ± 3 SD) (high BDI: .38%; low BDI: .61%).

For these analyses, ANOVAs on RT and discriminability were performed to assess putative group differences in task-unspecific variables (response speed, attentional processes, and task difficulty); as in prior signal-detection tasks, the main variable of interest was response bias (Henriques et al 1994; Tripp and Alsop 1999).

Finally, to directly assess the development of response bias after initial exposure to the differential reinforcement schedule in Block 1, difference scores between Blocks 1 and 2 and between Blocks 1 and 3 were computed. Because the difference score between Blocks 1 and 3 fully captured the overall development of response bias (ΔResponse Bias), this variable was entered within Pearson correlation analyses with self-report measures of mood and affect.

Results

Overall Effect of Task Manipulation (n = 62)

Accuracy. The two-way ANOVA with Block and Condition as factors revealed a main effect of Condition [F(1,61) = 29.50, p < .001 (partial η²: .33)] and a significant Block × Condition interaction [F(2,122) = 4.43, p < .020 (ε: .87; partial η²: .07)]. As expected, accuracy for the rich condition was higher than for the lean condition (Figure 2A), and this difference increased over time [linear contrast of the interaction: F(1,61) = 5.98, p < .020]. Post hoc Newman-Keuls tests clarified that accuracy for the rich condition was higher than for the lean condition in all three blocks (all p < .001). Also, the accuracy for the rich condition increased between Blocks 1 and 2 (p < .01) and between Blocks 1 and 3 (p < .005) but not between Blocks 2 and 3 (p > .25). Conversely, accuracy for the lean condition did not change over the blocks (all p > .60).

Reaction Time. The two-way ANOVA revealed significant effects of Block [F(2,122) = 7.98, p < .005 (partial η²: .12)], Condition [F(1,61) = 51.51, p < .001 (partial η²: .46)], and Block × Condition [F(2,122) = 5.94, p < .005 (partial η²: .09)]. As shown in Figure 2B, the rich condition led to significantly shorter RTs than the lean condition in all three blocks (Newman-Keuls: all p < .001). For both conditions, RTs were significantly shorter for Blocks 2 and 3 than for Block 1 (all p < .005), with no differences between Blocks 2 and 3.

Response Bias. The one-way ANOVA on response bias revealed a trend for Block [F(2,122) = 3.03, p = .06 (ε: .86)] due to a significant increase from Block 1 (.08 ± .22) to Block 2 (.14 ± .22; Newman-Keuls p < .050) and no further increase from Block 2 to Block 3 (.14 ± .19) (Figure 2C). Discriminability. The one-way ANOVA revealed no significant effects (all p > .10) (Figure 2D).

High (n = 15) Versus Low (n = 21) BDI Subjects

Self-Report Measures of Mood and Affect. High BDI subjects reported significantly higher dispositional anxiety (STAI) and negative affect (PANAS-NA) than low BDI subjects (Table 2).

Response Bias. The only effect emerging from the two-way ANOVA was the Group × Block interaction [F(2,68) = 3.27, p < .045 (ε: .59; partial η²: .09)]. As shown in Figure 3A, the interaction was due to increases in response bias for low BDI subjects from Block 1 (.09 ± .16) to Block 3 (.21 ± .14; Newman-Keuls p < .050), with no differences between Blocks 1 and 2 (p > .60) or between Blocks 2 and 3 (p > .20). Contrary to these findings, high BDI subjects failed to show any changes in response bias between the blocks (all p > .60; Figure 3A).

To directly assess whether high and low BDI subjects differed in their development of response bias, difference scores between Blocks 1 and 2 and between Blocks 1 and 3 were entered in a Group × Phase (Block 2 – Block 1; Block 3 – Block 1) ANOVA. The only significant effect was the interaction [F(1,34) = 4.45, p < .045; partial η²: .12] (Figure 3B). Compared with high BDI subjects, low BDI subjects had significantly higher increases in

A similar ANOVA considering only Blocks 2 and 3 confirmed that this effect was independent from response biases in Block 1, in which participants might have not been fully exposed to the differential reinforcement schedule; Group × Block interaction: F(1,34) = 4.45, p < .05. Furthermore, because high and low BDI subjects differed in trait anxiety and general negative affectivity, analyses were repeated entering separately the STAI and PANAS scores as covariates. The Group × Block interaction remained significant [STAI: F(1,18) = 6.71, p < .025; trait negative affect: F(1,31) = 3.58, p = .068].
response bias from Block 1 to Block 3 (Newman-Keuls \( p < .001 \)), with no group differences between Blocks 1 and 2 (\( p > .7 \)).

**Discriminability.** The two-way ANOVA revealed no significant effects.

**Reaction Time.** The three-way ANOVA revealed no significant effects involving Group.

**Control Analyses.** Because reinforcer ratio (the relative number of reinforcers received after a given correct response vs. another correct response) is the critical variable for producing a response bias (McCarthy and Davison 1979), it was important to rule out the possibility that the groups differed in the number of feedback received during the experiment. Separate \( t \) tests revealed that high and low BDI subjects received virtually identical reward feedback [rich: 88.40 \( \pm \) 1.77 vs. 88.10 \( \pm \) 2.88; lean: 29.40 \( \pm \) .63 vs. 29.52 \( \pm \) .68; rich/lean ratio: 3.01 \( \pm \) .09 vs. 2.99 \( \pm \) .13; all \( t(34) < .56, \) all \( p > .50 \)]. Finally, the two groups did not differ in the number of participants allocated to the sessions in which the short (high BDI: 8 of 15; low BDI: 11 of 21) or the long mouth was the condition more frequently rewarded.

**Correlations Between Self-Report Measures of Affect and Changes in Response Bias (n = 61)**

One participant had a \( \Delta \)Response Bias that was more than 4 SDs from the mean and thus was excluded from these analyses.\(^3\) Change in response bias from Block 1 to Block 3 (Block 3 – Block 1) was negatively correlated with BDI “melancholic” subscores (\( r = -.28, p < .055, n = 61 \)) assessed at Time 1. This change in \( \Delta \)Response Bias was not correlated with the Time 1 assessment of the Social Anhedonia (\( r = -.06, \) ns), Physical Anhedonia (\( r = -.06, \) ns), or general positive affectivity (PANAS-NA trait; \( r = -.09, \) ns) scores. Notably, \( \Delta \)Response Bias at Time 1 was negatively correlated with the total BDI score (\( r = -.46, p < .025, n = 25 \)) and BDI “melancholic” subscore (\( r = -.41, p < .05, n = 25 \); Figure 4) at Time 2. (These correlations remained significant when nonparametric Spearman Rank coefficients were used).

To test the specificity of these findings and the predictive value of response bias, two hierarchical regression analyses were run. In the first, we tested whether \( \Delta \)Response Bias at Time 1 predicted BDI “melancholic” subscores at Time 2 after controlling for BDI “melancholic” subscores at Time 1 (entered in the first step) and general negative affectivity (PANAS-NA trait).

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\(^3\)This subject had an intermediate BDI score (BDI = 14) and did not fulfill the inclusion criteria for the ANOVAs with the low and high BDI groups. Because the subject’s accuracy, RT, and response bias were within the mean \( \pm \) 2 SD, she was included in the ANOVAs with the entire sample (\( n = 62 \)).
When considering ΔResponse Bias at Time 1, 36.07% of the subjects had a negative change from Block 1 to Block 3, indicating that these participants were not responsive to the differential reinforcement schedule. As listed in Table 3, participants with negative ΔResponse Bias at Time 1 reported higher scores on various anhedonic measures at both Time 1 (BDI melancholic subscore, Chapman’s Physical Anhedonia scale) and Time 2 (BDI anhedonic subscore, BDI melancholic subscore, Anhedonic Depression subscale of the MASQ) and continued to have significantly lower ΔResponse Bias at Time 2.

Test-Retest Reliability

Twenty-five participants performed the signal-detection task a second time (38.28 ± 20.72 days later). As shown in Figure 5, ΔResponse Bias at Time 1 and Time 2 were significantly correlated ($r = .57, p < .004$).

Discussion

Rooted within a classic view that a stimulus is rewarding if it positively reinforces behavior (i.e., it reliably increases the likelihood of behavior [Hull 1943; Rescorla and Wagner 1972]), the goals of the present study were 1) to test the feasibility, convergent validity, and predictive validity of a laboratory-based measure of hedonic capacity based on a signal-detection approach; and 2) to test the hypothesis that elevated levels of depressive, particularly anhedonic, symptoms were associated with impaired reward responsiveness. Reward responsiveness was objectively operationalized as the participants’ propensity to modulate behavior as a function of prior exposure to reinforced stimuli. The following findings emerged. First, the task revealed robust behavioral differences, including a response bias, in favor of the stimulus reinforced more frequently. Second, unlike participants with low BDI scores, participants with elevated depressive symptoms failed to show increases in response bias toward the reinforcing stimulus, indicating reduced reward responsiveness. Importantly, high and low BDI participants did not differ in their discriminability and RT scores, or in the number of reward feedback received during the experiment, indicating that group differences in reward responsiveness were not due to differences in task difficulty, differential exposure to positive feedback, or different stimulus–outcome contingencies. Third, participants failing to develop a response bias throughout the

### Table 3. Self-Report Data for Participants with Positive (n = 39) Versus Negative (n = 22) ΔResponse Bias at Time 1

<table>
<thead>
<tr>
<th></th>
<th>ΔResponse Bias $&lt; 0$</th>
<th>ΔResponse Bias $&gt; 0$</th>
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<tbody>
<tr>
<td></td>
<td>Mean</td>
<td>SD</td>
</tr>
<tr>
<td>Time 1</td>
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</tr>
<tr>
<td>BDI</td>
<td>12.86</td>
<td>8.49</td>
</tr>
<tr>
<td>BDI Anhedonic Score $^a$</td>
<td>1.95</td>
<td>1.84</td>
</tr>
<tr>
<td>BDI Melancholic Score $^b$</td>
<td>4.09</td>
<td>2.74</td>
</tr>
<tr>
<td>Social Anhedonia $^c$</td>
<td>6.70</td>
<td>7.04</td>
</tr>
<tr>
<td>Physical Anhedonia $^d$</td>
<td>12.95</td>
<td>6.98</td>
</tr>
<tr>
<td>Time 2</td>
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<td></td>
</tr>
<tr>
<td>BDI</td>
<td>13.17</td>
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</tr>
<tr>
<td>BDI Anhedonic Score $^h$</td>
<td>2.22</td>
<td>2.11</td>
</tr>
<tr>
<td>BDI Melancholic Score $^e$</td>
<td>4.33</td>
<td>3.57</td>
</tr>
<tr>
<td>MASQ AD</td>
<td>70.22</td>
<td>12.65</td>
</tr>
<tr>
<td>ΔResponse Bias $^a$</td>
<td>−.01</td>
<td>.17</td>
</tr>
</tbody>
</table>

BDI, Beck Depression Inventory-II (Beck et al 1996); MASQ, Mood and Anxiety Symptom Questionnaire (Watson et al 1995); AD, Anhedonic Depression.

$^a$Difference score between response bias in Blocks 1 and 3 (ΔResponse Bias: Block 3 − Block 1).

$^h$Sum of BDI items associated with anhedonic symptoms: loss of pleasure (item 4), loss of interest (item 12), loss of energy (item 15), and loss of interest in sex (item 21).

$^e$Sum of BDI items associated with melancholic symptoms: loss of pleasure (item 4), guilty feelings (item 5), agitation (item 11), loss of interest (item 12), early morning awakening (item 16b), and loss of interest in sex (item 21).


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blocks (i.e., subjects with negative ΔResponse Bias) reported significantly higher scores on various self-report assessments of anhedonic and melancholic symptoms (e.g., loss of pleasure, energy, interest, and libido) in their daily activities, providing important convergence between subjective and objective measures of “anhedonia.” These findings are intriguing because BDI items assessing “loss of social interest,” “loss of appetite,” “loss of sexual interest and satisfaction,” and “work inhibition” best characterized depressed inpatients with extreme anhedonic features in an earlier clinical study (Fawcett et al 1983). Of note, participants failing to develop a response bias at Time 1 reported higher anhedonic symptoms at Time 2 and continued to have significantly lower ΔResponse Bias at Time 2 compared with participants developing a response bias. Furthermore, highlighting the specificity of these findings, hierarchical regression analyses clarified that ΔResponse Bias at Time 1 predicted “melancholic” symptoms at Time 2 even after controlling for initial “melancholic” symptoms and general negative affectivity. Therefore, the lower the response bias development in the signal-detection task, the higher the degree of melancholic, including anhedonic, symptoms at the day of the experiment as well as 1 month later. Finally, a test–retest analysis suggested a moderately strong stability, within a 1-month period (r = .57), of the propensity to develop a response bias.

No reliable correlations emerged when considering the Chapman’s anhedonia scales. One possible explanation for this null finding is that these scales might have predictive validity primarily for schizophrenia spectrum disorders (Kwapil 1998) and thus might be less sensitive for assessing dysphoria or depression. In line with this view, a recent study found that a state measure of anhedonia that was correlated with depressive symptoms was unrelated to the Physical Anhedonia scores (Herbener and Harrow 2002). An alternative, not mutually exclusive, explanation could be that the anhedonic items included in the BDI (e.g., loss of libido, loss of pleasure) might be more closely related to our operationalization of hedonic capacity. More work is needed to clarify this issue.

Approach-Related System Abnormalities in Depression: Replication and Extension

Collectively, the present findings of blunted responsiveness to reinforcing stimuli in subjects with elevated depressive symptoms agree with prior empirical evidence and theoretic propositions that have suggested a critical role of an abnormal approach-related system in depression (Costello 1972; Davidson 1992; Depue and Iacono 1989; Loas 1996; Watson and Clark 1984). Particularly, our findings replicate and extend earlier reports in depression of blunted hedonic responses, which, as in the present study, were assessed with laboratory-based measures. Using a progressive ratio schedule with increasing reinforcement, Hughes et al (1985) showed in a small study with six melancholic depressed participants that the number of responses needed to obtain money increased in the three treatment responders but not in the three nonresponders. Using a task assessing the effect of incentives on performance during a verbal memory task, Henriques and coworkers (Henriques et al 1994; Henriques and Davidson 2000) reported that dysphoric and clinically depressed subjects displayed a more conservative bias during the reward condition; depressed subjects showed a weaker tendency to report that a word had been presented before, a strategy that in control subjects led to maximization of earning. Together with reports of low levels of self-reinforcement (Gotlib 1982), underestimation of frequency of positive reinforcement (Buchwald 1977), diminished response to pleasant cues (Berenzbaum and Oltmanns 1992; Sloan et al 2001), and decreased attention to positive facial expression (Suslow et al 2001), these findings suggest that the use of laboratory-based measures might offer a powerful tool for objectively assessing hedonic capacity in depression. The present findings specifically clarify that impairments in the extent to which behavior can be modulated by reinforcement history might be a potential mechanism underlying an abnormal approach-related system in depression. Because reinforcing stimuli that increase the frequency of specific behavioral responses (Hull 1943; Rescorla and Wagner 1972) and play a critical role in the formation of associations between salient cues and internal rewarding events (Spanagel and Weiss 1999), lowered responsiveness to reinforcers might lead to diminished engagement in pleasurable activities and decreased motivational drive to pursue future rewards (Depue and Iacono 1989; Kasch et al 2002). Although longitudinal studies will be required to determine whether impaired reward responsiveness might represent a vulnerability factor to depression, we speculate that subtle impairments in hedonic responsiveness might predispose individuals to lower engagement with the environment. In line with this conjecture, studies using self-report measures of affect have reported that low behavioral activation system levels (particularly, reward responsiveness) negatively predicted outcome 8 months later (Kasch et al 2002). Similarly, a deficit in positive affect represented a risk factor for the development and/or maintenance of depressive symptoms in both community (Lonigan et al 1999) and inpatient (Joiner and Lonigan 2000) samples of children and adolescents. Moreover, in a large population-based study, anhedonia at study entry was associated with poor outcome 12 months later (Spijker et al 2001). The present finding that low reward responsiveness predicted anhedonic symptoms 1 month later is consistent with and extends these prior findings, which were based on self-report measures.

Limitations

The limitations of this study should be acknowledged. First, although the BDI is a reliable and well-established measure of depressive symptomatology (e.g., Beck et al 1988), this self-report measure is not indicative of the full clinical range of...
depression. Consequently, although the present report indicates that elevated levels of depressive, particularly anhedonic, symptoms were associated with decreased reward responsiveness, it is unknown whether these findings will extend to a clinical sample. Moreover, some of the statistically significant results emerging from the present study were relatively weak. For example, the correlation between change in response bias and BDI melancholic subscore was \( r = -0.28 \). Accordingly, future studies will be required to evaluate the reliability, specificity, and predictive validity of the present signal-detection approach to assess hedonic capacity in clinically depressed patients. Additional studies will also be needed to test the specificity of the present findings with respect to clinical nosology and to assess whether the proposed laboratory-based approach might be useful for distinguishing among different psychopathologies. One could speculate that schizophrenic patients, particularly those with a deficit syndrome, might show impaired hedonic capacity and reduced ability to discriminate between stimuli (reflected in abnormal discriminability scores). Furthermore, individuals with anxiety disorders might show normative hedonic capacity but increased response bias toward punishment-related cues, in line with data indicating attentional biases toward threat-related cues in anxiety disorders (Mineka et al. 2003).

**Implication for Neuroscience Research on Depression**

In conclusion, a signal-detection task was used to provide an objective assessment of participants’ propensity to modulate behavior as a function of reward-related cues. Participants reporting elevated depressive symptoms failed to modulate their behavior despite receiving identical monetary rewards, indicating that they were less sensitive or reactive to pleasurable stimuli (i.e., more anhedonic). Critically, participants with diminished response bias reported relatively lower pleasure/interest in their daily activities even 1 month later, providing important convergent and predictive validity to the “anhedonia” construct used in the present study. These findings confirm that deficits in an approach-related system might be a cardinal feature of depression, and warrant further development of laboratory-based measures of symptoms in clinical populations.

As recently reviewed by Hasler et al. (2004), one of the main challenges in biological psychiatry will be to improve the phenotypic definition of depression so that a better understanding of the genetic and neurobiological underpinnings of this debilitating disease can be achieved. Dissecting the depressive phenotypes into key components yielding narrow definitions of psychopathologic endophenotypes is of utmost priority in reaching this goal (Hasler et al. 2004; Hyman 2002; Leboyer et al. 1998); endophenotypes are indeed expected to provide a “means for identifying the ‘downstream’ traits or facets of clinical phenotypes, as well as the ‘upstream’ consequences of genes” (Gottesman and Gould 2003 [p. 637]). It was in this spirit that the present study was designed, and research is currently underway in our laboratory adopting this approach for identifying more homogenous subtypes of depression and assessing neural underpinnings underlying impaired hedonic capacity in depression.

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